

**Title:** Traffic-related air pollution near busy roads: the East Bay Children's Respiratory Health Study

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## Abstract

Recent studies, primarily in Europe, have reported associations between respiratory symptoms and residential proximity to traffic; however, few have measured traffic pollutants or provided information about local air quality. We conducted a school-based cross-sectional study in the San Francisco Bay Area in 2001. Information on current bronchitis symptoms and asthma, home environment, and demographics were obtained by parental questionnaire (n=1,109). Concentrations of traffic pollutants (particulate matter (PM<sub>10</sub>, PM<sub>2.5</sub>), black carbon (BC), and nitrogen oxides (NO<sub>x</sub> and NO<sub>2</sub>)) were measured at ten school sites during several seasons. Although pollutant concentrations were relatively low, we observed differences in concentrations between schools nearby versus those more distant (or upwind) from major roads. Using a two-stage multiple logistic regression model, we found associations between respiratory symptoms and traffic-related pollutants. Among those living at their current residence for at least one year, the adjusted odds ratios (OR) for asthma in relation to an interquartile difference in NO<sub>x</sub> were OR = 1.07; (95% confidence interval, 1.00-1.14). Thus, we found spatial variability in traffic pollutants and associated differences in respiratory symptoms in a region with good air quality. Our findings support the hypothesis that traffic-related pollution is associated with respiratory symptoms in children.

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Suggested MeSH: Air Pollution, Vehicle Emissions, Asthma, Bronchitis, Epidemiology

## **Introduction**

Numerous epidemiological studies have documented adverse effects of air pollution on health. (1) The majority of these population-based studies have used pollutant concentrations measured at central monitoring sites to estimate exposures and have not, in general, considered local spatial variability in pollutant levels. However, motor vehicle emissions, the principal source of ambient air pollution in most urban areas, are likely to vary substantially within a given community, and researchers have begun to document differences in traffic-related pollutants on a neighborhood scale.(2,3)

Recently, a number of epidemiological studies have reported associations between residential proximity to busy roads and a variety of adverse respiratory health outcomes in children, including respiratory symptoms, asthma exacerbations, and decrements in lung function.(4-10) In some reports, truck traffic has been more strongly associated with these adverse outcomes than total vehicular traffic. (6,7,10) (11)

Most studies have used metrics of proximity to traffic as surrogates of exposure to traffic pollution (e.g., residential proximity to major roads, traffic volume at the nearest road, or modeled levels of traffic pollution). Few have measured pollutant concentrations as part of the exposure assessment or provided information on local air quality. (7,10,12) (11) The majority of studies have been conducted in Europe and Japan, where fleet composition (diesel vs. gasoline), emissions factors, fuel specifications, land use, and population distributions near busy roads differ from those in the U.S. Regional and micro-environmental concentrations of particulate matter may be higher in European

cities compared with many parts of the United States.(13) Therefore, it is important to evaluate the extent to which proximity to traffic may be associated with health impacts in the United States. Previous studies in the U.S. were conducted in areas of Southern California and the Northeast with significant local air quality problems; both used metrics of proximity to traffic, not measured pollutant concentrations.(8) (14)

The objective of this study was to explore associations between respiratory symptoms and exposures to traffic-related air pollutants among children living and attending schools near busy roads in an urban area with high traffic density, but good regional air quality. Some of the results of these studies have been previously reported in the form of abstracts.(15)

## **Methods: (word count 552)**

### *Study design and health assessment*

We conducted a school-based cross-sectional study in the San Francisco metropolitan area (Alameda County, California) in 2001. The study area was comprised of ten neighborhoods that span a busy traffic corridor. School sites were selected to represent a range of locations upwind and downwind of major roads (Figure 1).

In Spring 2001, we enrolled children (Grades 3-5) in participating classes (n=64) using methods similar to those used in other school-based studies. (16)(17,18) We obtained information on health outcomes (bronchitis symptoms in the past 12 months and

physician-confirmed asthma in the past 12 months), demographics, home environmental factors and activity factors using parental questionnaires (English and Spanish).(16-18) For additional information on the study design and health assessment, see the online supplement.

The study protocol was approved by the Committee for the Protection of Human Subjects, California Health and Human Services Agency.

*Air Pollution from Traffic:*

We measured concentrations of traffic pollutants (particulate matter (PM<sub>10</sub>, PM<sub>2.5</sub>), black carbon (BC), and nitrogen oxides (NO<sub>x</sub> and NO<sub>2</sub>)) at the school sites. PM<sub>10</sub> and PM<sub>2.5</sub> mass concentrations were measured using filter-based samples, while BC concentrations were determined on the PM<sub>10</sub> filter samples using an established light attenuation method that we validated for fiberfilm filters.(19) (20) NO<sub>x</sub> and NO<sub>2</sub> concentrations were determined with passive diffusion samplers (Ogawa, Inc., USA). Nitric oxide (NO) concentrations were calculated as the difference between NO<sub>x</sub> and NO<sub>2</sub>.

Pollutant monitoring was conducted simultaneously at all school sites for 11 one-week intervals in the spring (March - June) and for eight weeks in the fall (September-November) of 2001. NO<sub>x</sub> and NO<sub>2</sub> were sampled during all weeks at each school. PM<sub>10</sub> and PM<sub>2.5</sub> and the BC concentrations were not measured every week. Study-averaged air pollution concentrations were calculated at each school by first normalizing the data to account for occasional missing values. Additional details are described in the

online data supplement and elsewhere.(21)

In preliminary analyses, we also used school location in relation to prevailing winds and proximity to busy roads as an additional traffic metric. .

### Data Analysis

We examined associations between pollutants and health outcomes using a two-stage hierarchical modeling strategy. This method has been used in other epidemiological studies of air pollution when pollutants were measured at the group level.(22) (18) In our study, the exposure groups are represented by the neighborhood schools. In the first stage, we initially identified potential confounders (demographic, host, or home environmental variables) associated with health outcomes in this dataset. We then performed exploratory stepwise logistic regressions to develop a model in which individual-level characteristics best predicted the odds of each health outcome. Explanatory variables that remained significant at  $p < 0.15$  were retained in the model. We then fit a logistic regression model that included an indicator variable for each school in addition to the individual-level covariates.

In the second stage, the adjusted school-level logits or prevalence rates determined in the first stage were regressed on the school-specific ambient pollutant concentrations. In this manner, we obtained the log odds ratios relating asthma or bronchitis symptoms to air pollution, after adjusting for individual-level risk factors.

We calculated adjusted odds ratio(s) (ORs) for a change in measured pollutant

concentration equal to the interquartile ranges (IQR) of the pollutant distributions.

Analyses were conducted using SAS version 8.2 for Windows (Cary, NC) and STATA version 8 (College Park, TX).

## **Results**

We distributed 1,574 questionnaires in 64 participating classrooms in the 10 schools. Three children were excluded because their parents spoke neither English nor Spanish. Among the remaining students, there was a response rate of 70.7 % (1,111/1,571). Participation rates across schools ranged from 61 to 83%. Approximately 30% completed the questionnaire in Spanish. Two children with reported cystic fibrosis were excluded from the analysis. The final analysis sample consisted of 1109 questionnaires.

Table 1 summarizes the participants' demographic characteristics, prevalence of selected personal and home environmental characteristics and respiratory health outcomes. Our study population was racially diverse. About 30% of households had incomes below the federal poverty line. Fourteen percent of the parental respondents reported having been told by a doctor that their child had asthma in the preceding 12 months. This represents a measure of period prevalence of asthma and would include some incident cases. Twelve percent of children had bronchitis symptoms in the past year. Of those reporting bronchitis symptoms in the past 12 months, 43% also reported having asthma. Using a slightly different definition of asthma (physician-diagnosed ever, and asthma symptoms, including wheezing, in the past 12 months), 11 % of our study

population had current asthma.

The estimated pollutant concentrations at the schools are summarized in Table 2. Concentrations of several pollutants (i.e., BC, NO<sub>x</sub>, NO and, to a lesser extent, NO<sub>2</sub>) were higher at schools located within 300 meters downwind of a freeway compared to those at schools upwind or further from major traffic sources. There was less variation in PM<sub>2.5</sub> and PM<sub>10</sub>. Concentrations of BC, NO<sub>x</sub>, and NO were highly correlated ( $r^2 \sim 0.9$  for each inter-pollutant correlation). The study average PM<sub>2.5</sub> (12 µg/m<sup>3</sup>) was similar to the annual average concentration of PM<sub>2.5</sub> at the central monitoring station, located approximately 15 kilometers south of the study area. NO<sub>x</sub> and NO<sub>2</sub> measurements at the school sites away from traffic were similar to levels measured at the regional site. (21).

Table 3 summarizes the results of the two-stage hierarchical logistic regression models of the odds of asthma and bronchitis symptoms in the previous year in relation to six different pollutants, each examined in separate regressions. Results are shown for all subjects, for long-term residents only (one year or longer at the current address), and for the latter group stratified by gender. In addition to the traffic metric, explanatory variables retained in all the final models for asthma and bronchitis included chest illness before age two, household mold/moisture, and pests observed in the home in the preceding 12 months. The final models for asthma also included maternal history of asthma. Addition of other potential confounders such as race/ethnicity, socioeconomic variables, maternal smoking during pregnancy, current smoker in the home, air conditioning, and gas stove use, yielded similar pollutant effect estimates.

For the full sample, associations were observed between both asthma in the past 12 months and bronchitis symptoms in the past 12 months and the pollutants, especially NO<sub>x</sub>, NO, and BC. The effect estimates for PM<sub>2.5</sub> and PM<sub>10</sub> were smaller, which may have been due in part to the smaller concentration ranges among the 10 sites for these pollutants. No multi-pollutant models were evaluated because of the high inter-pollutant correlations. Restricting the analysis based on duration of residence (i.e., at least one year at current residence) tended to increase the effect estimates slightly in relation to asthma, especially when the sample was restricted to girls. Stratification by duration of residence or gender did not change the results for bronchitis. Results were similar when non-normalized pollution values were used (data not shown).

We conducted additional sensitivity analyses, including: (1) dropping the one school that was an outlier with respect to the proportion of Hispanic students (89% versus 21% -53% at other schools); (2) using a different definition for current asthma; and (3) stratifying bronchitis by a reported history of asthma. When the “outlier” school was dropped, the magnitude of the odds ratios for bronchitis did not change much, but the confidence intervals were wider. In the asthma analyses, dropping the outlier school resulted in similar or slightly greater effect estimates. Applying different questionnaire-based asthma definitions showed little change but slightly larger confidence intervals. After stratifying students by whether they also “ever” had asthma, the results suggested that those with a history of asthma were driving the results for bronchitis, but the sample size became too small to make clear inferences.

Figures 2 and 3 depict the associations between BC and bronchitis and asthma.

## **Discussion**

To our knowledge, this is the first epidemiological study in the United States to evaluate relationships between measured traffic-related pollutants and respiratory symptoms. For children residing at their current address for at least one year, we found modest but significant increases in the odds of bronchitis symptoms and physician-diagnosed asthma in neighborhoods with higher concentrations of traffic pollutants. These results are consistent with previous reports of positive associations between proximity to traffic and various respiratory outcomes. (4-10) (11),(7,10-12) Furthermore, our findings were observed in a region with relatively clean air (low concentrations of ozone and particulate matter). (See online supplement for details) Although previous epidemiological studies in the United States exploring chronic respiratory effects of air pollution in children have shown inconsistent results, this might be due in part to exposure misclassification as these studies used air quality measurements conducted at single fixed-site monitors in each city. (17,18,22,23)

Our findings were robust to multiple sensitivity analyses using different questionnaire-based definitions of current asthma and wheezing in the past 12 months. The slight increase in effect estimates for associations between asthma after restricting the analysis to those with longer duration at current residence may be due to a reduction in exposure measurement error. Our study population was very mobile (23% had moved in the preceding 12 months, only 32% had lived at the same address since before age

two).

We considered whether there might be bias due to non-response or self-reporting. We saw no significant difference in proportions of questionnaires returned in Spanish versus English by school, but there was a modest inverse correlation between pollution concentrations measured at each school and response rate. However, the response rate for individual classrooms within each school varied as well and appeared to depend on the willingness of teachers to encourage participation. Dropping the school closest to a freeway (which also had the highest measured pollutant concentrations, a high percentage of Hispanic students, and the lowest response rate) did not change the effect estimates for bronchitis and increased the estimates for asthma. This would suggest that knowledge of potential high traffic exposure probably did not affect parental reporting of the children's respiratory histories. This study was not undertaken in response to public concerns about traffic, nor, at the time the study was conducted, was there much local interest in potential health hazards of proximity to traffic. Therefore, reporting and non-response biases were unlikely to have unduly influenced our results.

We found increased association with asthma (but not bronchitis) with exposure to traffic air pollutants for girls who had lived at their current addresses at least one year compared with boys (Table 3). Several investigators have also reported greater traffic-associated effect estimates for girls versus boys.(7,8,10,24,25) Previous air pollution studies examining the gender-specific effects of air pollution on lung function and lung function growth have been mixed (26,27) The reasons for the observations in our study are unclear

and deserve attention in future studies.

### Exposures

We found spatial variability in exposure due specifically to roads with heavy traffic within a relatively small geographic area for BC, NO<sub>x</sub>, NO, and to a lesser extent, NO<sub>2</sub>. There was less variation in PM<sub>2.5</sub> across schools; this is consistent with previous observations that PM<sub>2.5</sub> is more likely to reflect regional air quality. (2) The higher effect estimates with BC, NO<sub>x</sub>, and NO compared with NO<sub>2</sub> and PM<sub>2.5</sub>, suggest that primary or fresh traffic emissions may play an etiologic role in these relationships. While NO<sub>x</sub>, NO, and BC may serve as indicators of exposure to traffic-related pollutant mixtures, they may also act as etiologic agents themselves.(28)

We found that downwind direction was an important determinant of increased exposure to traffic pollutants, and that a simple traffic indicator (school location downwind and <300m from a major road) gave estimates of odds ratios similar to or greater than pollutant measurements in preliminary analyses using a one-stage model (data not shown). Within a geographic area with flat terrain and low-rise buildings, the direction of wind in relation to the traffic source is the most important weather parameter. Other parameters important in air dispersion of traffic pollutants (e.g., atmospheric stability, wind speed, and surface topography) would be relatively similar at the different school sites.

A simple single-stage logistic model using pollutant measurements also yielded positive associations between pollutants and symptoms with a much larger effect estimate

and smaller confidence intervals.

We assumed that traffic-related pollutants measured at the neighborhood schools would be a good proxy for the children's overall exposure to such pollutants. Children attending the schools in this study generally lived within walking distance and did not use school buses. Therefore, pollutant concentrations in the children's neighborhoods probably tracked those at their schools. The most plausible exposure error in an urban setting would be that subjects who attend schools with very high traffic exposures from a nearby freeway would tend to have similar or lower home exposures whereas children with low school exposures would tend to live in homes with similar or only slightly higher traffic exposures. This pattern of measurement error would tend to underestimate the association between exposure and outcome. (29)

Alternatively, repeated daily exposures for 6-8 hours during the school year may themselves represent biologically important influences on some children's respiratory health, analogous to occupational exposures for susceptible adults. In a recent study of proximity to traffic and respiratory health, Janssen et al. found that effect estimates using based on the school-to-highway distance were comparable or greater than those based on residence-to-highway distance. (11)

The average measurements at each school were used to estimate long-term average traffic air pollutant concentrations. We measured pollutants at each of the 10 sites concurrently (to avoid concerns of week-to week variability) in two different periods that reflect the major seasonal wind patterns for the area. We found that the rank order

(relative values) of the schools did not vary from week to week or season to season, supporting the validity of this approach. Additionally, the NO<sub>x</sub> and NO<sub>2</sub> concentrations at schools upwind or further from high traffic roads were similar to NO<sub>x</sub> and NO<sub>2</sub> concentrations measured at the closest fixed-site monitor (21). Although there may have been some changes in the absolute traffic volume on major roads in recent years, the principal traffic patterns in the area have not changed. Thus, the relative values (rank order) of the site-specific pollutant concentrations measured in our study are likely to be representative of those in recent years.

The cross-sectional nature of our study design is a further limitation on causal inference, but we observed the same or modest increase in effect estimates for current asthma and bronchitis when we restricted our analysis to those who had lived at their present address for at least a year. Most studies on proximity to traffic and respiratory symptoms have been cross-sectional, and further longitudinal studies are needed to elucidate the role of traffic-related air pollution in the development and exacerbation of asthma and other respiratory symptoms.

Another limitation was that the exposures were assigned at the group level (n=10); however, the multi-level analysis allows adjustment for individual confounders in the first stage of analysis. Moreover, in this respect this study is comparable with other epidemiological investigations (e.g., the Harvard Six Cities Study and the Children's Health Study in Southern California (n=12 communities). Another recent cross-sectional study of traffic-related air pollution and respiratory symptoms included 13 schools.(22)

(18,23)

We also lacked information on indoor measurements of traffic-related pollutants. However, recent studies have found high correlations between personal exposures to NO<sub>2</sub> and traffic parameters. (30) Others have found that indoor concentrations and exposure to soot (particulate matter from diesel exhaust) is highly correlated with outdoor levels.(2)

#### Other covariates.

Maternal asthma, household mold/moisture, pests, and chest illness before age two were important explanatory variables in the final model for current asthma, consistent with previous studies (31).(32,33) We explored whether current levels of traffic pollution could modify the risk of current asthma symptoms depending on past history of chest illness; however, there was not sufficient power to explore interactions based on early medical history. Race/ethnicity and indicators of socioeconomic status were not important predictors of health outcomes in our study. This may be due, in part, to our study design (i.e., the schools were selected to have relatively similar measures of SES).

We did not find associations between exposure to environmental tobacco smoke and current asthma; the results of previous cross-sectional studies in school-aged children have been mixed. (34) The prevalence of current household smokers in our study was small, however, limiting study power. It is possible that there is some under-reporting of household smoking (7 % in our study vs. 19% statewide).(35) Alternatively, a substantial portion of our study population were less acculturated Hispanics (30% of parents

responded in Spanish), and only 3.6% of Hispanic households reported a history of maternal smoking. Other investigators have also observed very low smoking rates (less than 5%) among less acculturated Hispanics (B. Eskenazi, personal communication).(36) If under-reporting does exist, it is possible that residual confounding might have affected our estimates of pollutant/respiratory health outcome relationships. However, the addition to the regression model of variables correlated with exposure to environmental tobacco smoke (ETS) (e.g., SES and race-ethnicity) did not change the pollutant effect estimates, suggesting that significant confounding by ETS was not likely.

In summary, we found associations between traffic-related pollutants and asthma and bronchitis symptoms in the past 12 months in a highly urbanized region of the United States with good regional air quality, where local air pollution is dominated by vehicular sources. Although the cross-sectional study design, exposure assignment at the group level, small geographic area, and possible unmeasured covariates, may limit the generalizability of the study, our findings are consistent with previous investigations in Europe and the United States. (11,14,37) In addition, our results underscore the limitations of using central air monitoring stations for assigning population exposures. Concentrations of air toxics such as diesel exhaust particles or surrogates such as BC or soot should be more widely monitored. Measurement of personal exposures to traffic pollutants is not feasible in large population-based studies; the use of geographic modeling approaches to estimate exposures for individuals may be a good alternative.(38) Future studies that can better characterize exposures to traffic pollutants and their sources (i.e., diesel vs. gasoline engines) will be important to better understand the public health

impacts of motor vehicle emissions.

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Figure Legends.

Figure 1: East Bay Children's Respiratory Health Study area. The study region is to the east and across the bay from the city of San Francisco.

Figure 2. Adjusted School-Specific Bronchitis Prevalence Rates Versus Black Carbon, Long-term Residents

Figure 3. Adjusted School-Specific Asthma Prevalence Rates Versus Black Carbon, Long Term Residents

Table1. Demographic, family, and home characteristics of the East Bay Children’s Respiratory Health Study respondents.

Characteristics	Subjects attending schools		
	All subjects (N = 1109) %	Near and Downwind of major roads (4 schools, N =402) %	Far or Upwind of major roads (6 schools, N = 707) %
<b>Gender</b>			
Female	52.6	51.8	53.1
<b>Race/Ethnicity</b>			
White	12.6	11.0	13.5
Black, African American	11.1	7.0	13.4
Hispanic	43.5	47.6	41.2
Asian	14.0	15.5	13.1
Other/Multiracial	18.9	18.8	18.9
<b>SES indicators</b>			
Household at/below Federal poverty level	31.3	31.8	31.0
Parent's education: high school or less *	48.7	51.4	47.1

**Family history**

Biological mother with asthma	12.2	9.5	13.7
Maternal smoking during pregnancy	10.3	7.8	11.7

**Home indoor environment**

Smoker in the household, since child's birth	17.9	13.1	20.6
Smoker in the household, current	7.2	3.2	9.5
Furry pet	37.3	36.0	38.1
Pests, past 12 months	63.1	65.4	61.8
Gas stove	63.1	63.6	62.9
Indicator of mold/mildew presence, past 12 months	44.6	43.5	45.3

**Health outcomes**

Chest illness before age 2	23.3	18.8	25.9
Asthma, past 12 months	14.0	13.9	14.1
Bronchitis, past 12 months	12.1	13.2	11.5

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\*Parent responding to the questionnaire

Table 2: Nearby traffic sources and average pollutant concentrations at ten schools\*

School	Major Traffic Source <sup>†</sup>	AADT (#/day)	Distance <sup>‡</sup> <300 m (m)	downwind	PM <sub>10</sub> (µg/m <sup>3</sup> )	PM <sub>2.5</sub> (µg/m <sup>3</sup> )	BC (µg/m <sup>3</sup> )	NO <sub>x</sub> ppb	NO <sub>2</sub> ppb	NO <sup>§</sup> ppb
1	No				30	12	0.7	42	22	19
2	Yes	90,000	230	Yes	29	13	0.9	55	24	31
3	Yes	210,000	360	No	32	12	0.8	49	21	29
4	No				30	12	0.8	41	19	22
5	Yes	210,000	130	Yes	30	12	0.9	62	26	36
6	No				30	12	0.7	39	21	17
7	No				29	11	0.7	33	20	11
8	Yes	130,000	350	No	29	12	0.7	45	23	21
9	Yes	210,000	200	Yes	30	12	0.9	57	26	31
10 <sup>11</sup>	Yes	190,000	60	Yes	32	15	1.1	69	31	38
Study average					30	12	0.8	49	23	25

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\*Estimated average pollutant concentration at each school based on normalized concentrations (see text). Monitoring was conducted for 11 weeks in the spring (March - June) and eight weeks in the fall (September-November) 2001). The number of weeks underlying our estimates of chronic exposure varied for each measured pollutant: NO<sub>x</sub> (18), NO<sub>2</sub> (19), BC (11), PM<sub>2.5</sub> (10) and PM<sub>10</sub> (9).

†Includes roads with annual average daily traffic (AADT) above 50,000 vehicles per day located within 1000 m of school.

AADT estimate provided by CA Dept of Transportation (Cal Trans).

‡Distances were estimated using geographic information systems. Latitude and longitude of the monitors were determined using a global-positioning system (GPS) device (Garmin GPS II). In some cases, distances were estimated using aerial photographs or measured using a distance wheel.

§NO = NO<sub>x</sub>-NO<sub>2</sub>

<sup>11</sup>There are also a shopping center and parking lot abutting the school grounds to the south and a freeway off ramp <50 m to the northwest.

Table 3. Odds ratios (95% Confidence Interval) of respiratory illness by school-based ambient air pollutant concentrations using two-stage model.

Exposure	<u>All subjects</u>		<u>LTR subjects<sup>‡</sup></u>		<u>LTR-Females<sup>‡</sup></u>		<u>LTR-Males<sup>‡</sup></u>	
	(N=1109)		(N=871)		(N=462)		(N=403)	
	OR	CI	OR	CI	OR	CI	OR	CI
<b>Bronchitis*</b>	N=93/797		N=79/635		N=38/341		N=41/291	
NO <sub>x</sub>	1.05	(1.01, 1.08)	1.06	(1.03, 1.09)	1.07	(1.03, 1.11)	1.03	(0.98, 1.09)
NO <sub>2</sub>	1.02	(0.99, 1.06)	1.03	(1.00, 1.06)	1.04	(1.01, 1.08)	1.02	(0.98, 1.06)
NO	1.05	(1.02, 1.09)	1.06	(1.03, 1.09)	1.07	(1.03, 1.11)	1.04	(0.98, 1.10)
PM <sub>10</sub>	1.03	(0.99, 1.07)	1.02	(0.98, 1.07)	1.04	(1.01, 1.09)	1.01	(0.95, 1.06)
PM <sub>2.5</sub>	1.02	(1.00, 1.05)	1.03	(1.01, 1.05)	1.04	(1.02, 1.05)	1.02	(0.99, 1.05)
BC	1.04	(1.00, 1.08)	1.05	(1.01, 1.08)	1.06	(1.02, 1.10)	1.03	(0.98, 1.08)
<b>Asthma<sup>†</sup></b>	N=101/705		N=78/562		N=42/297		N=36/263	
NO <sub>x</sub>	1.04	(0.97, 1.11)	1.07	(1.00, 1.14)	1.17	(1.06, 1.29)	1.02	(0.93, 1.11)
NO <sub>2</sub>	1.02	(0.97, 1.07)	1.04	(0.98, 1.10)	1.09	(1.03, 1.15)	1.00	(0.94, 1.07)
NO	1.05	(0.98, 1.12)	1.08	(1.00, 1.15)	1.19	(1.03, 1.36)	1.02	(0.94, 1.12)
PM <sub>10</sub>	1.02	(0.96, 1.09)	1.04	(0.97, 1.12)	1.09	(0.92, 1.29)	1.02	(0.94, 1.10)

	PM <sub>2.5</sub>	1.00	(0.97, 1.04)	1.01	(0.97, 1.06)	1.06	(0.99, 1.15)	0.99	(0.95, 1.04)
	BC	1.02	(0.96, 1.09)	1.05	(0.99, 1.13)	1.12	(0.95, 1.33)	1.00	(0.93, 1.09)
<b>Asthma (no outlier, school 5)<sup>†</sup></b>		N=96/641		N = 73/507		N = 38/271		N = 35/233	
	NO <sub>x</sub>	1.08	(1.00, 1.17)	1.10	(1.00, 1.20)	1.14	(1.02, 1.28)	1.07	(0.96, 1.19)
	NO <sub>2</sub>	1.06	(0.99, 1.13)	1.07	(0.98, 1.17)	1.09	(0.97, 1.22)	1.05	(0.96, 1.16)
	NO	1.08	(1.00, 1.17)	1.09	(1.00, 1.19)	1.14	(1.03, 1.26)	1.07	(0.96, 1.18)
	PM <sub>10</sub>	1.06	(0.97, 1.16)	1.08	(0.98, 1.19)	1.09	(0.96, 1.24)	1.08	(0.97, 1.19)
	PM <sub>2.5</sub>	1.04	(0.96, 1.12)	1.03	(0.94, 1.13)	1.03	(0.91, 1.17)	1.03	(0.94, 1.14)
	BC	1.07	(0.98, 1.17)	1.09	(0.99, 1.19)	1.14	(1.02, 1.27)	1.06	(0.95, 1.18)

\* First stage model adjusted for: child's respiratory illness before age 2; pests; indicator of mold presence

†First stage model adjusted for: child's respiratory illness before age 2; pests; indicator of mold presence; maternal history of asthma

‡LTR = long term resident (current address for 1 year or more)

Odds ratios are calculated per interquartile range (IQR) of average pollutant concentration as follows:

NO<sub>x</sub> = 14.9 ppb, NO<sub>2</sub> = 3.6 ppb, NO = 11.6 ppb, PM<sub>10</sub> = 1.4 µg/m<sup>3</sup>, PM<sub>2.5</sub> = 0.7 µg/m<sup>3</sup>, BC = 0.15 µg/m<sup>3</sup>

Figure 1.

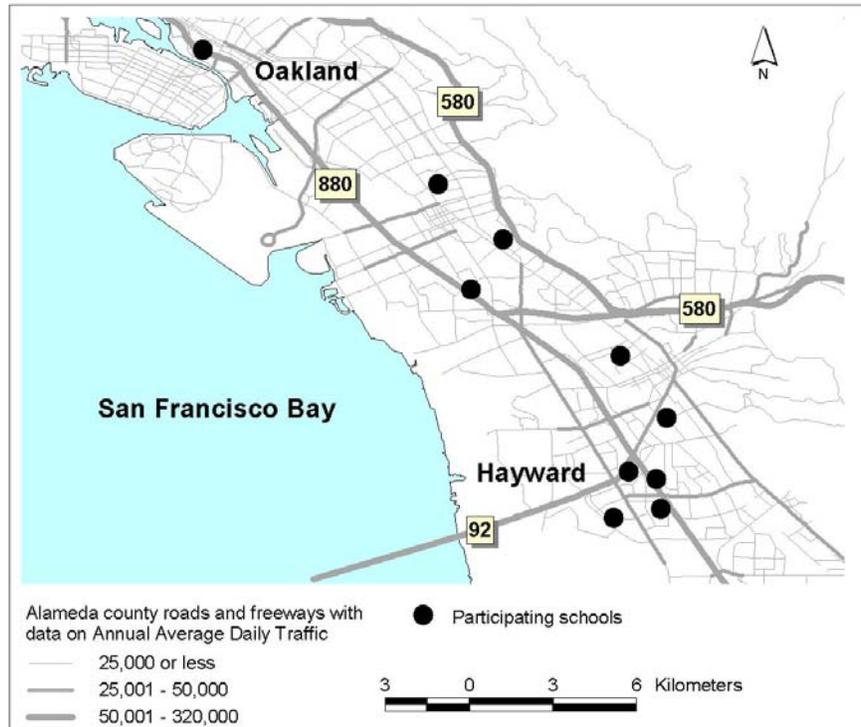


Figure 2

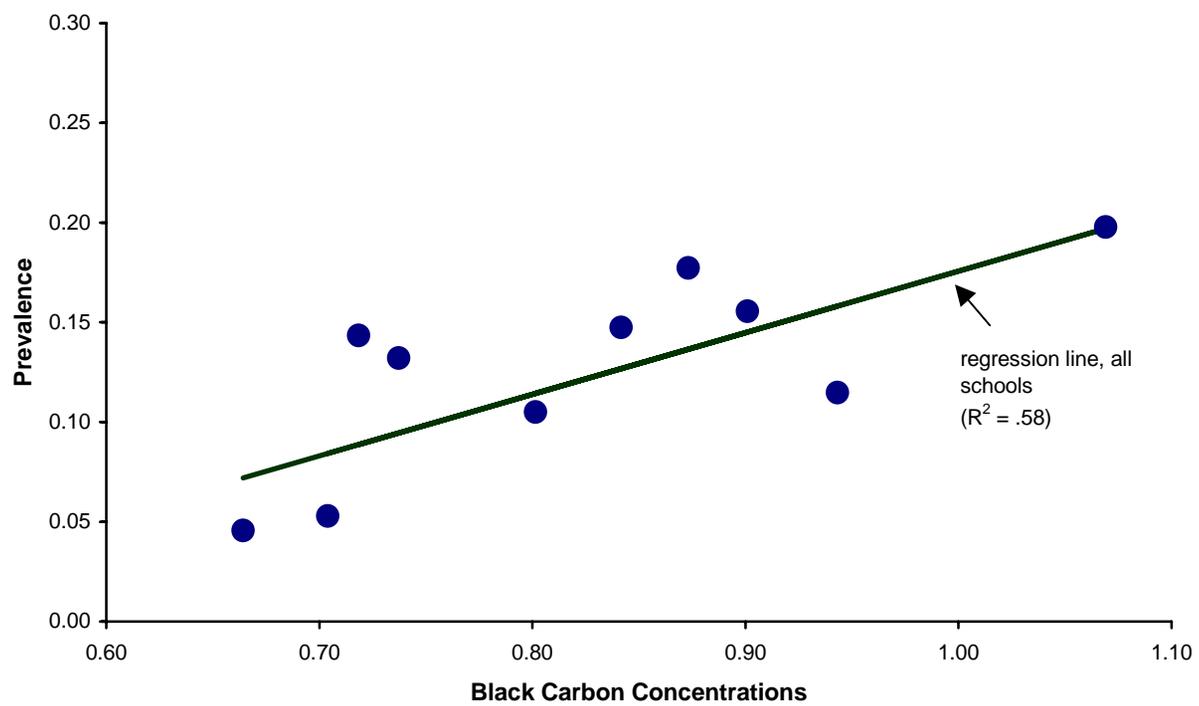
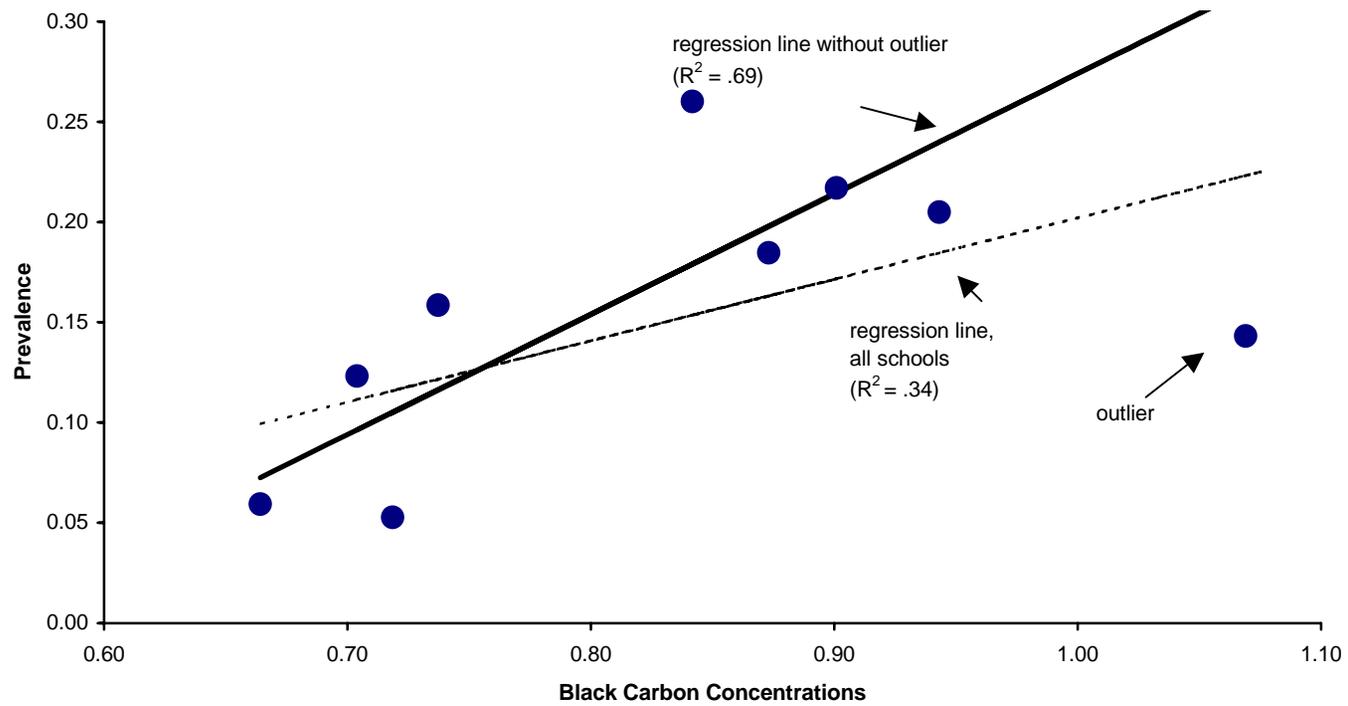


Figure 3.



**Title:** Traffic-related air pollution near busy roads: the East Bay Children's Respiratory Health Study

**Authors**

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**Online Data Supplement**

## **Methods :**

### *Study design and study location*

The study area (approximately 21 km x 5 km) is across the bay from San Francisco, and includes 10 neighborhood elementary schools that span a busy traffic corridor. The topography is relatively flat and the urban landscape consists primarily of widely spaced, low-level buildings. Traffic congestion in the San Francisco-Oakland area ranks second in the United States,(39) but regional air quality is generally good due to ocean breezes from the southwest or west. At the nearby air quality monitoring station (Fremont), 15 km southeast of the study area, ozone concentrations rarely exceed the national or state standards (0-1days/ year above federal ozone standards for past 3 years). Annual-average concentrations of PM<sub>10</sub> and PM<sub>2.5</sub> at the Fremont station were 23.4 and 12.2 µg/m<sup>3</sup>, respectively in 2001. These values are below the PM<sub>10</sub> and PM<sub>2.5</sub> National Ambient Air Quality Standards (50 µg/m<sup>3</sup> and 15 µg/m<sup>3</sup>, respectively) and just above the more stringent California standards ( 20 µg/m<sup>3</sup> and 12 µg/m<sup>3</sup>, respectively) .

### **School site selection**

We used a public school database from the California Department of Education (CDE) containing statewide information on school locations, enrollment, and demographics (URL: <http://www.cde.ca.gov>). A traffic database was obtained from the California Department of Transportation (CalTrans) Highway Performance Monitoring System (HPMS, 1997). Traffic files included statewide freeway and major road network

geography and counts of the annual average daily traffic (AADT) on those roadways.

We utilized Geographic Information Systems (GIS) software to integrate, display, and analyze data from these disparate data sources (ArcView 3.2, ESRI, Redlands, CA).

We selected elementary schools for possible recruitment such that the school locations would represent a range of distances upwind and downwind from major roads. Because concentrations of traffic pollutants decrease to background levels within 300 m downwind (40,41) and because several European studies have found adverse health effects in children residing or attending schools in close proximity to roads with AADT of 25,000 or more vehicles (4,9), we used these parameters to guide our selection process.

Briefly, to identify possible school sites for recruitment, school addresses were geo-coded and overlaid with the CalTrans road network. Candidate schools “near” a major road were identified based on a geo-coded location less than 350 m from any road with AADT  $\geq 90,000$  vehicles. Schools “farther” from local or heavy traffic were also identified using GIS (no local road with AADT  $\geq 20,000$  vehicles/day within 300 m and no major highway or freeway within 750 m). We identified eight schools that were “near” freeways and twenty-one schools that were “farther” from local or heavy traffic in the proposed study area. From this list, we sought to recruit schools that were demographically similar and that were likely to have a range of exposure to traffic pollutants based on proximity to major roads and prevailing wind directions.

Ten schools that reflected locations with a range of distances upwind and downwind from major roads were selected. There were no major non-roadway sources of air pollution in

the neighborhoods (CA Air Resources Board, private communication).

Population:

In Spring, 2001, we invited all children in grades 3-5 in participating classes (n=64) to join the study. Packets containing a health questionnaire in English or Spanish, an informed consent form, and study information were distributed in the classroom, completed by parents, and returned to the teacher. An option was available to complete the questionnaire by telephone with trained research staff. Up to three reminder notices were sent to non-responders. We gave a donation towards educational materials to classrooms and provided nominal compensation to parents completing the questionnaire. The study protocol was approved by the Committee for the Protection of Human Subjects, California Health and Human Services Agency.

Health Assessment:

Questions on respiratory symptoms and illnesses were modeled after those used in previous studies on air pollution and children's health.(16) (17,18) Subjects were defined as having current asthma if their parents selected asthma in response to the question, "During the previous 12 months, did a doctor say that your child had any of the following chest illnesses?" (Choices were pneumonia, asthma, reactive airway disease, and other chest illness.) Current bronchitis was defined as: (1) a positive response to the question: "During the past 12 months, did your child have bronchitis?" or (2) a report of cough and chest congestion or phlegm lasting at least three consecutive months of the past 12.

We also asked about personal factors potentially associated with asthma or bronchitis, including demographic variables (race/ethnicity and measures of socioeconomic status (SES) (e.g., parents' education, income, health insurance)); host factors (age, gender, chest illness before age two, maternal smoking during pregnancy, maternal or paternal history of asthma, premature birth, breastfeeding); home environmental factors (current smoker in the home, gas stove, evidence of mold, pets, problem with pests in the past year, carpet in the bedroom; and activity factors (sports participation, commute patterns to school). Pests included cockroaches, mice, rats, termites, spiders, or ants. Those with "chest illness before the age of two" were reported to have seen a doctor before age two for asthma, bronchiolitis, RSV, croup, reactive airway disease, or pneumonia.

Indicator of "mold" was a Yes to any of the following: history of residential water damage, visible mold/yildew, water condensation, or mold/musty smell in the past 12 months. In sensitivity analyses we used a slightly different definition of asthma (physician-diagnosed ever, and asthma symptoms, including wheezing, in the past 12 months), and duration of residence (at least one year at current address Yes/No).

#### Air Pollution from Traffic:

To estimate the children's overall exposure to traffic pollution, we measured concentrations of traffic-related pollutants at the schools. We also used school location with respect to prevailing winds and proximity to busy roads as a separate exposure metric.

(1) Traffic-related pollutant measurements: We determined  $PM_{10}$  and  $PM_{2.5}$  mass

concentrations from weeklong filter samples (Pallflex fiberfilm, Pall/Gelman) collected with small sampling pumps (Airchek 2000, SKC, Inc) and size-selective inlets (PEM, MSP Corp). BC concentrations were determined on the PM<sub>10</sub> filter samples using an established light attenuation method that we validated for the fiberfilm filters.(19) (20) Concentrations of NO<sub>x</sub> and NO<sub>2</sub> were determined with passive diffusion samplers (Ogawa, Inc., USA) deployed over a one-week sampling period. The NO concentration was estimated as the difference between NO<sub>x</sub> and NO<sub>2</sub>. Additional details of the monitoring protocol are described elsewhere.(21)

We estimated chronic exposures to traffic pollutants by measuring pollutant concentrations simultaneously at all schools over several month. Specifically, monitoring was conducted for 11 weeks in the spring (March - June) and eight weeks in the fall (September-November) 2001). NO<sub>x</sub> and NO<sub>2</sub> were sampled during all periods at each school.(21) PM<sub>10</sub> and PM<sub>2.5</sub> were not measured every week.

Study-averaged air pollution concentrations were calculated at each school by first normalizing the data to account for occasional missing data as follows: Because of missing samples and week-to-week variability at each site, we obtained average site-specific pollutant concentrations as follows: For each week and pollutant, we calculated the ratio of each school's measured pollutant level relative to the mean across all schools sampled that week. Then, for each school, the mean ratio over the entire study period was calculated. The mean ratios are normalized concentrations and represent the relative exposures across school locations. To obtain an estimate of the average pollutant concentration at each school, we multiplied the mean of the school-specific normalized

concentration by the study-average pollutant concentration (averaged across all study sites over the entire study period). This normalization process allowed for an aggregation of data for all sampling periods, including those with an incomplete data set. Only weeks with valid data from at least six schools were included in this process.

(2) Additional traffic metric: Schools were assigned a category for wind direction (upwind or downwind of a freeway) and a category for proximity to traffic (<300 m or >300 m from a major traffic source) as well as the combination of the two (i.e., downwind and close versus all others).

### Data Analysis

To explore the extent to which additional covariates would have an effect on the pollutant estimate, we added pollutant values or school location indicators to the model, one at a time and then examined the effect of removing each covariate on the pollutant coefficient(s); if removal of a covariate from the model changed the magnitude of the pollutant coefficient by less than 10%, the covariate was dropped.

In a separate sensitivity analysis, we also initially included variables found to be predictive of respiratory symptoms in other studies even if they did not satisfy the p-value criterion. Other several sensitivity analyses included examining the model using other definitions of asthma and after stratifying the outcome of current bronchitis by asthma status. We also restricted the sample to the subgroup of children who had lived at the given address for more than one year (long-term residents) and performed stratified analysis by gender, both for the full sample and long-term residents. Additionally, we

re-ran the models after dropping one school that was an outlier in terms of race/ethnicity distribution.

Adjusted odds ratio(s) (ORs) were calculated for an interquartile range (IQR) of measured pollutant concentrations (i.e., the odds ratio for a given health outcome given a pollutant concentration at the 75 percentile of the distribution relative to that at the 25<sup>th</sup> percentile). All analyses were conducted using SAS version 8.2 for Windows and STATA version 8.

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